

Homework 5 Solutions

1)

a) Here is the stata output

```
. xi: logistic case i.agegp tob alc
i.agegp      _Iagegp_1-6      (naturally coded; _Iagegp_1 omitted)

Logistic regression              Number of obs   =       976
                                LR chi2(7)        =       294.47
                                Prob > chi2        =       0.0000
Log likelihood = -347.74036      Pseudo R2     =       0.2975
```

case	Odds Ratio	Std. Err.	z	P> z	[95% Conf. Interval]	
_Iagegp_2	5.851322	6.433949	1.61	0.108	.6781047	50.49069
_Iagegp_3	34.90621	36.8794	3.36	0.001	4.401332	276.8351
_Iagegp_4	59.30331	62.43568	3.88	0.000	7.53212	466.918
_Iagegp_5	101.4181	107.7435	4.35	0.000	12.64226	813.5914
_Iagegp_6	105.0871	116.9321	4.18	0.000	11.86871	930.4551
tob	1.041306	.0082032	5.14	0.000	1.025352	1.057509
alc	1.026306	.0026633	10.01	0.000	1.021099	1.031539

The interpretation of the odds ratio for tob is that 1.0413 is the odds of oesophageal cancer associated with a one gram/day increase in tobacco consumption, after adjusting for age and alcohol consumption. The OR for alc indicates that the odds of oesophageal cancer increases by a factor of 1.0263 for each one gram/day increase in alcohol consumption, after adjusting for age and tobacco use.

b) Stata output for adding TOB²:

```
. xi: logistic case i.agegp tob alc tob2
i.agegp      _Iagegp_1-6      (naturally coded; _Iagegp_1 omitted)

Logistic regression              Number of obs   =       976
                                LR chi2(8)        =       296.01
                                Prob > chi2        =       0.0000
Log likelihood = -346.96762      Pseudo R2     =       0.2990
```

case	Odds Ratio	Std. Err.	z	P> z	[95% Conf. Interval]	
_Iagegp_2	5.318698	5.844352	1.52	0.128	.6172746	45.82814
_Iagegp_3	31.73698	33.4744	3.28	0.001	4.015739	250.822
_Iagegp_4	54.35493	57.10336	3.80	0.000	6.934209	426.0701
_Iagegp_5	91.37431	96.9344	4.26	0.000	11.42424	730.8376
_Iagegp_6	96.9282	107.657	4.12	0.000	10.99083	854.8101
tob	1.066909	.0226092	3.06	0.002	1.023503	1.112155
alc	1.026023	.0026677	9.88	0.000	1.020808	1.031265
tob2	.9994097	.0004758	-1.24	0.215	.9984777	1.000343

The Wald test for TOB² has p-value equal to .21 suggesting that a quadratic term is not necessary for tobacco.

Stata output for adding ALC²:

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```
. xi: logistic case i.agegp tob alc alc2
i.agegp      _Iagegp_1-6      (naturally coded; _Iagegp_1 omitted)

Logistic regression                                Number of obs   =       976
                                                    LR chi2(8)      =       294.84
                                                    Prob > chi2     =       0.0000
Log likelihood = -347.55224                        Pseudo R2       =       0.2978
```

case	Odds Ratio	Std. Err.	z	P> z	[95% Conf. Interval]	
_Iagegp_2	5.924965	6.543814	1.61	0.107	.6801085	51.61708
_Iagegp_3	36.12426	38.36408	3.38	0.001	4.506302	289.5861
_Iagegp_4	61.16769	64.71265	3.89	0.000	7.691116	486.4686
_Iagegp_5	104.5539	111.6036	4.36	0.000	12.90461	847.1014
_Iagegp_6	105.9373	118.2657	4.18	0.000	11.87925	944.7326
tob	1.041279	.0082051	5.13	0.000	1.025321	1.057486
alc	1.021612	.0081205	2.69	0.007	1.005819	1.037652
alc2	1.00003	.0000504	0.60	0.547	.9999316	1.000129

The Wald test for ALC^2 has p-value equal to .55 suggesting that a quadratic term is not necessary for alcohol.

- c) Since both individuals are 60 years old, I can ignore the age term in the model in calculating the OR (a reasonable estimate of the RR). After refitting the model in (a) (to make it the current model) I can do

```
. lincom 80*alc + 15*tob
( 1) 15 tob + 80 alc = 0
```

case	Odds Ratio	Std. Err.	z	P> z	[95% Conf. Interval]	
(1)	14.64964	3.535333	11.12	0.000	9.128729	23.50952

Given that oesophageal cancer risk is rare, I estimate that the *risk* of oesophageal cancer for the drunken smoker is 14.6 times the risk of the clean-livin' teetotaler.

2)

a) An equivalent null hypothesis in terms of odds ratios is

$H_0: 1 = OR_2 = OR_3 = OR_4$ (i.e., take $H_0: odds_1 = odds_2 = odds_3 = odds_4$ and divide by $odds_1$ to get H_0 as OR's.)

where $OR_i = p_i/(1-p_i)/[p_1/(1-p_1)]$ and $p_i = \text{Pr}(\text{disease} | \text{alcohol level } i)$. Each odds is compared to the lowest alcohol group. You could, of course, choose any of the alcohol groups as the reference category and get an equivalent null hypothesis.

The stata output is

```
. tabodds y alc [freq=count]
```

alc	cases	controls	odds	[95% Conf. Interval]	
<40g/day	29	386	0.07513	0.05151	0.10957
40-79g/~y	75	280	0.26786	0.20760	0.34560
80-119g~y	51	87	0.58621	0.41489	0.82826
120+g/day	45	22	2.04545	1.22843	3.40587

```
Test of homogeneity (equal odds): chi2(3) = 158.79
Pr>chi2 = 0.0000
```

```
Score test for trend of odds: chi2(1) = 152.97
Pr>chi2 = 0.0000
```

We conclude there is a highly significant ($p < .001$) trend in the odds of disease with level of alcohol exposure.

b) The stata output for a logistic regression is

```
. logit y alc [freq=count]
```

```
Logit estimates                Number of obs   =      975
                               LR chi2(1)         =    144.64
                               Prob > chi2         =     0.0000
Log likelihood = -422.4246      Pseudo R2       =     0.1462
```

y	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
alc	1.046772	.0935048	11.19	0.000	.8635064	1.230038
_cons	-3.530124	.2279715	-15.48	0.000	-3.976939	-3.083308

The test of $H_0: \beta_1 = 0$ is rejected with $p < .001$. We estimate that the log odds of disease increases by 1.05 (95% CI 0.86 – 1.23) for each categorical increase of alcohol consumption.

c) Here's the stata output

```
. xi: logit y i.alc [freq=count]
i.alc          _Ialc_1-4          (naturally coded; _Ialc_1 omitted)

Logit estimates                                Number of obs   =          975
                                                LR chi2(3)      =         146.50
                                                Prob > chi2     =          0.0000
Log likelihood = -421.49545                    Pseudo R2      =          0.1481
```

y	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
_Ialc_2	1.27124	.232332	5.47	0.000	.8158777	1.726602
_Ialc_3	2.054459	.2611044	7.87	0.000	1.542704	2.566214
_Ialc_4	3.304162	.3236511	10.21	0.000	2.669817	3.938506
_cons	-2.588542	.1925445	-13.44	0.000	-2.965922	-2.211161

Point estimates and confidence intervals for each level of alcohol use are given. For instance, the coefficient for `_Ialc_3` is 2.05 (95% CI 1.54 – 2.57) indicating that the logodds for disease associated with 80 – 119 gms/day alcohol consumption is 2.05 greater than the logodds for disease associated with 0 - 39 gms/day alcohol consumption.

d) Here is the relevant stata output

```
. generate alc3=0
. replace alc3=1 if alc==3
. generate alc4=0
. replace alc4=1 if alc==4
. xi: logit y alc alc3 alc4 [freq=count]

Logit estimates                                Number of obs   =          975
                                                LR chi2(3)      =         146.50
                                                Prob > chi2     =          0.0000
Log likelihood = -421.49545                    Pseudo R2      =          0.1481
```

y	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
alc	1.27124	.232332	5.47	0.000	.8158777	1.726602
alc3	-.488021	.3685031	-1.32	0.185	-1.210274	.2342318
alc4	-.5095586	.6067233	-0.84	0.401	-1.698714	.6795972
_cons	-3.859782	.406446	-9.50	0.000	-4.656401	-3.063162

		Predicted probabilities			
	likelihood	$\pi(\text{ALC}=1)$	$\pi(\text{ALC}=2)$	$\pi(\text{ALC}=3)$	$\pi(\text{ALC}=4)$
Model (c)	-421.495	.0699	.2112	.3695	.6716
Model (d)	-421.495	.0699	.2112	.3695	.6716

I conclude that models (c) and (d) are equivalent.

Since it is clear that the model

$$\text{logit}[\pi(X)] = \beta_0 + \beta_1 \text{ALC}$$

is nested within the model

$$\text{logit}[\pi(X)] = \beta_0 + \beta_1 \text{ALC} + \beta_2 \text{ALC}(3) + \beta_3 \text{ALC}(4)$$

and we have shown that the models

$$\text{logit}[\pi(X)] = \beta_0 + \beta_1 \text{ALC} + \beta_2 \text{ALC}(3) + \beta_3 \text{ALC}(4)$$

$$\text{logit}[\pi(X)] = \beta_0 + \beta_1 \text{ALC}(2) + \beta_2 \text{ALC}(3) + \beta_3 \text{ALC}(4)$$

are equivalent, that means that

$$\text{logit}[\pi(X)] = \beta_0 + \beta_1 \text{ALC}$$

is nested within the model

$$\text{logit}[\pi(X)] = \beta_0 + \beta_1 \text{ALC}(2) + \beta_2 \text{ALC}(3) + \beta_3 \text{ALC}(4).$$

- e) The likelihood ratio chi square is equal to 1.86 (2 df) and the p-value is 0.39. Thus we would conclude that the logistic linear model in part b gives an adequate fit to these data.
f) The stata output is

```
. xi: logit y alc i.age [freq=count]
i.age          _Iage_1-6      (naturally coded; _Iage_1 omitted)
```

Logit estimates	Number of obs	=	975
	LR chi2(6)	=	255.91
	Prob > chi2	=	0.0000
Log likelihood = -366.78769	Pseudo R2	=	0.2586

y	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
alc	1.093726	.1032947	10.59	0.000	.8912717 1.29618
_Iage_2	1.575683	1.07541	1.47	0.143	-.532081 3.683448
_Iage_3	3.304031	1.031874	3.20	0.001	1.281595 5.326467
_Iage_4	3.826275	1.027804	3.72	0.000	1.811816 5.840734
_Iage_5	4.214935	1.033778	4.08	0.000	2.188767 6.241102
_Iage_6	4.308531	1.083401	3.98	0.000	2.185104 6.431957
_cons	-6.983926	1.050337	-6.65	0.000	-9.042548 -4.925304

The odds ratio for alc in this model is $\exp(1.09) = 2.98$ (95% CI = 2.44 – 3.66). We interpret this as the adjusted (for age) relative increase in the odds of disease for each categorical increase in alcohol consumption.

- g) The likelihood ratio test comparing a model with age categories versus a model with age as a linear term has a chi-square statistic equal to 17.77 with 4 df. The p value is 0.0014 so we would conclude that the linear model for age is not adequate.
h) I would reason like this: each of the categories of alcohol consumption in our current model represents about a 40g/day increase in alcohol consumption. The coefficient in this model is 1.046. So if we had a model with actual grams of alcohol consumption, I might expect that the value of the coefficient would be about $1.046/40 = 0.026$.