# Penalized loss functions for Bayesian model comparison Martyn Plummer, *Biostatistics* (2008)

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## Model Selection in Bayesian Models

### Bayes Factor:

- The formal solution
- Unstable with diffuse prior; undefined with improper priors

### Cross-validation:

- Which model is most useful?
- Judge model by out-of-sample prediction

#### Posterior-Predictive approach:

- Does this model give data like what I observed?
- Simulate from posterior and compare to original data

## Model Selection in Bayesian Models

#### **Deviance Information Criterion (DIC)**

 $DIC = \overline{D} + p_D \approx$  "Model Fit" + "Model Complexity"

- Proposed by Spiegelhalter et al. (2002)
- Theoretical foundations are controversial
  - No clear generalization outside of exponential families
  - Sensitive to parameterization
- "Experience with DIC to date suggests that it works remarkably well" –Banerjee *et al.* (2004)

### Can we develop a formal justification for DIC?

## Loss Functions for Model Selection

Suppose we have a set of data  $\mathbf{Y} = (Y_1, \dots, Y_n)$  $Y_i \sim p(\cdot | \boldsymbol{\theta}) \quad \boldsymbol{\theta} \sim \pi(\cdot)$ 

What is a suitable loss function for model comparison?

- Decision Theory suggests using scoring rules, which are functions of p(·)
- Maximized when  $p(\cdot)$  is true data generating density

Deviance:

$$D(\theta) = -2\log\{p(\mathbf{Y}|\theta)\}$$

## Computing the Deviance

Ideally we have two data sets:

- Training data Z
- Test data Y
- $\blacktriangleright p(\mathbf{Y}|\boldsymbol{\theta}, \mathbf{Z}) = p(\mathbf{Y}|\boldsymbol{\theta})$

Plug-in Deviance:

$$L^{p}(\mathbf{Y}, \mathbf{Z}) = -2 \log \left[ p\{\mathbf{Y} | \mathrm{E}(\boldsymbol{\theta} | \mathbf{Z})\} \right]$$

Expected Deviance:

$$L^{e}(\mathbf{Y}, \mathbf{Z}) = -2 \int \log\{p(\mathbf{Y}|m{ heta})\} \pi(m{ heta}|\mathbf{Z}) \, dm{ heta}$$

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## Penalized Loss

Typically, we only have one set of data **Y**. Can we use  $L(\mathbf{Y}, \mathbf{Y})$ ?

Yes, but we're being optimistic by using the data to both estimate the posterior of  $\theta$  and as our test data

Add a penalty term to loss function:

 $L(\mathbf{Y}, \mathbf{Y}) + p_{opt}$ 

## Penalized Loss

We can split our loss function into contributions from each  $Y_i$ 

$$L(\mathbf{Y},\mathbf{Y}) = \sum_{i=1}^{n} L(Y_i,\mathbf{Y})$$

Compare  $L(Y_i, \mathbf{Y})$  to cross-validation loss  $L(Y_i, \mathbf{Y}_{-i})$  to estimate how optimistic we are being.

$$p_{opt_i} = \mathrm{E}\Big[L(Y_i, \mathbf{Y}_{-i}) - L(Y_i, \mathbf{Y})\Big|\mathbf{Y}_{-i}\Big]$$

The penalized loss function is now

$$L(\mathbf{Y},\mathbf{Y}) + \sum_{i} p_{opt_i}$$

### Penalized Loss and DIC

We will see that

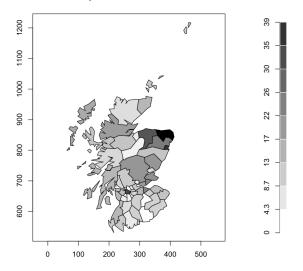
$$DIC \approx L^p(\mathbf{Y}, \mathbf{Y}) + \sum_i p_{opt_i}$$

...but only when the effective number of parameters is small relative to the number of observations.

When this is not true, DIC will under-penalize complex models

## Penalized Loss and DIC in Disease Mapping

Lip Cancer in Scotland



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Penalized Loss and DIC in Disease Mapping

 $Y_i \sim \text{Poisson}(\mu_i) \quad \log(\mu_i) = \alpha_0 + \gamma_i + \delta_i + \log(E_i)$ 

 $Y_i$  – lip cancer cases in county i

 $E_i$  – expected counts of lip cancer in county *i* 

 $\alpha_0$  – fixed effect

 $\gamma_i$  – uncorrelated random effects

 $\delta_i$  – spatially correlated random effects

Model	DIC	Penalized Loss
Fixed Effect Only	1.0	1.1
Uncorrelated	43.5	570.5
Spatial	31	163.9
Uncorrelated + Spatial	31.6	166.4

Table: Estimated penalties for model complexity for Scottish lip cancer data

## Looking Ahead

Looking more closely at  $DIC \approx L^{p}(\mathbf{Y}, \mathbf{Y}) + \sum_{i} p_{opt_{i}}$ 

#### Application to Scotland Lip Cancer Data

Can we apply penalized loss functions to other settings?

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